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Linking amygdala persistence to real-world emotional experience and psychological well-being

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Competing Interests

R.J.D. serves on the board of directors for the non-profit organization Healthy Minds Innovations. The other authors report no perceived or real conflicts of interest.

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Abstract

Neural dynamics in response to affective stimuli are linked to momentary emotional experiences. The amygdala, in particular, is involved in subjective emotional experience and assigning value to neutral stimuli. Because amygdala activity persistence following aversive events varies across individuals, some may more readily evaluate subsequent neutral stimuli than others. This may lead to more frequent and long-lasting momentary emotional experiences, which may also be linked to self-evaluative measures of psychological well-being (PWB). Despite extant links between daily affect and PWB, few studies have directly explored the links between amygdala persistence, daily affective experience, and PWB. To that end, we examined data from 52 human adults (67% female) in the Midlife in the US (MIDUS) study who completed measures of PWB, daily affect, and functional MRI (fMRI). During fMRI, participants viewed affective images followed by a neutral facial expression, permitting quantification of individual differences in the similarity of amygdala representations of affective stimuli and neutral facial expressions that follow. Using representational similarity analysis, neural persistence to aversive stimuli was operationalized as similarity between the amygdala activation patterns while encoding negative images and the neutral facial expressions shown afterwards. Individuals demonstrating less persistent activation patterns in the left amygdala to aversive stimuli reported more frequent positive and less frequent negative affect in daily life. Further, daily positive affect served as an indirect link between left amygdala persistence and PWB. These results clarify important connections between individual differences in brain function, daily experiences of affect, and well-being.

Significance

At the intersection of affective neuroscience and psychology, researchers have aimed to understand how individual differences in the neural processing of affective events map onto to real-world emotional experiences and evaluations of well-being. Using a longitudinal dataset from 52 adults in the Midlife in the US (MIDUS) study, we provide an integrative model of affective functioning: less amygdala persistence following negative images predicts greater positive affect in daily life, which in turn predicts greater psychological well-being seven years later. Thus, day-to-day experiences of positive affect comprise a promising intermediate step that links individual differences in neural dynamics to complex judgements of psychological well-being.

Introduction

Psychological well-being (PWB) captures one's perceived self-acceptance, positive relations with others, autonomy, environmental mastery, purpose in life, and personal growth (Ryff, 1989; Ryff & Keyes, 1995). Relatively more enduring, trait-like judgements of PWB are linked to relatively more transient positive and negative emotional states in the real-world (Burns & Ma, 2015; Rush et al., 2019). Specifically, higher PWB is associated with greater positive affect (PA; Burns & Ma, 2015) and lower negative affect (NA; Rush et al., 2019) in daily life. A key question is what neurobiological processes give rise to these subjective affective experiences and judgements. One promising source of these individual differences may be the time course of amygdala activity. (Sander et al., 2003).

The amygdala is necessary for threat detection and threat conditioning—the assigning of value to neutral stimuli (LeDoux, 1996). The amygdala is also implicated in a range of affective processes, including general detection of salience (Davis & Whalen, 2001; Kober et al., 2008; Sander et al., 2003), facial processing (Todorov et al., 2008), and experiencing fear (LeDoux, 2000). In particular, the amygdala supports negative appraisals of putatively neutral stimuli when they are preceded by an unrelated, aversive stimulus (Lapate et al., 2016). Individual differences in amygdala response to fearful faces predict how negatively a subsequent, unrelated neutral stimulus is appraised (Lapate et al., 2016; Lapate et al., 2017). This biasing effect may be due to persistence of amygdala activity following emotionally evocative stimuli that ‘spills over’ into the encoding of the subsequent neutral stimulus (Grupe et al., 2018; Tambini et al., 2017).

Functional MRI (fMRI) research suggests that individual differences in the degree of amygdala activity persistence after aversive events is associated with individual differences in affective style. Schuyler et al. (2014) found that more persistent amygdala

activity following negative images predicted higher levels of neuroticism. Additionally, greater amygdala persistence following negative images was linked to lower likeability ratings of subsequent neutral faces. Similarly, greater amygdala persistence following negative words has been found in patients with major depressive disorder (Siegle et al., 2002). These studies demonstrate that amygdala persistence is linked to individual differences in affective style.

However, the existing literature has relied on univariate amygdala analyses that cannot account for the multivariate representation of affective stimuli. Assessing persistence by averaging activity over the amygdala removes fine-grained, spatial activity patterns that carry detailed information about stimulus properties. In contrast, multivariate approaches, such as representational similarity analysis (RSA), preserve the voxel-wise pattern of activation and compare the neural representations of different types of stimuli. Although RSA was initially used to examine similarity in visual perception (Kriegeskorte et al., 2008; Kriegeskorte & Bandettini, 2007), it has been applied to identify similarity in affective processing (Brooks & Freeman, 2018; Ölander et al., 2017). Given that we rarely encounter a single stimulus in isolation, the dynamics of amygdala representation across stimuli are paramount to understanding individual differences in well-being. Therefore, by capturing the commonalities in the voxel-wise pattern of amygdala activity among affective stimuli, compared with the pattern of activity among subsequent neutral facial expressions, RSA could yield a more accurate and ecologically valid metric of amygdala persistence.

To that end, the current study used RSA to test whether individual differences in amygdala persistence following affective images was linked with daily affective experience and PWB. We analyzed a MIDUS subsample that completed a PWB questionnaire, daily telephone interviews of affective experience, and an fMRI scan. We hypothesized that those with less amygdala persistence following negative visual stimuli,

reflected in less similarity between voxel-wise neural representations of negative stimuli and neutral facial expressions that follow negative stimuli, would report higher daily PA and less daily NA. In line with previous work, we further hypothesized that higher reports of daily PA and less NA would correspond to greater PWB. Lastly, using a mediational framework, we tested whether daily affect provided a pathway by which amygdala persistence was linked with PWB.

Materials and Methods

Participant Characteristics

Data were collected from 2004-2009 as a part of the MIDUS-II Longitudinal Study of Health and Well-being (<http://www.midus.wisc.edu/>), which recruited a national sample (ages 35-85) through random digit dialing. Participants first completed self-report questionnaires ($n = 4963$); a random subsample then completed a series of eight daily consecutive telephone interviews to assess daily affect ($n = 2,022$), and a further subgroup able and willing to travel to the Midwest included neuroscience assessments, including an fMRI scan ($n = 72$). Two participants did not complete all five functional runs of the MRI task; eight were removed for excessive motion (sudden spikes, gradual drift, or both, that exceeded 2 millimeters). Nine of the remaining participants did not complete the daily diary study. Finally, one participant was excluded from the analysis for reporting levels of daily negative affect greater than three standard deviations from the group mean. This resulted in 52 participants that completed MRI, daily diary, and self-reported PWB for the analyses (39-76 years old ($M = 57.74$, $SD = 10.5$); 67% female; 69% Caucasian, 29% African American, 2% Native American or Alaskan Native). There were five sets of twins in the sample. Lastly, 31 of these subjects repeated the self-report questionnaires pertaining to PWB again 7 years later ($SD = 1.6$, range = 5 – 9 years).

On average, the first measurement of PWB was completed 729 days, or 24.3 months, prior to the daily diary procedure (SD = 430 days, range = 136 – 1533 days) and 1230 days, or 41 months, prior to the fMRI scan (SD = 372 days, range = 348 – 1669 days). Thirty-one participants completed the daily diaries prior to the fMRI session by an average of 999 days (SD = 347 days, range = 408 – 1491 days). The other twenty-one participants completed the fMRI scan prior to the daily diary procedure by an average of 231 days (SD = 59, range = 161 – 386 days). Because significant time passed between each measurement and time lag varied by participants, each relevant time lag is included as a covariate in the analysis. Therefore, significant relationships from the analyses are likely to reflect stable, meaningful connections that withstand long and variable measurement gaps. Data and documentation for MIDUS 1, MIDUS 2, MIDUS 3, and all MIDUS projects are available to the public for downloading at the Inter-university Consortium for Political and Social Research (ICPSR) website (<https://www.icpsr.umich.edu/web/ICPSR/series/203>).

Experimental Design and Statistical analyses

Psychological Well-being

Participants completed a 42-item measure of PWB (Ryff, 1989; Ryff & Keyes, 1995). This measure contains subscales for six domains of PWB, including self-acceptance, positive relations with others, autonomy, environmental mastery, purpose in life, and personal growth. Each subscale is comprised of 7 items that are rated on a 1-7 scale ('strongly agree' to 'strongly disagree'). Items that indicate greater PWB are reverse coded such that higher overall scores reflect higher PWB. The total score for the measure is calculated by averaging the score from each of the subscales. This measure of PWB has been shown to be reliable over the span of 6 weeks (test-retest reliability for

the six scales > 0.8 ; Ryff, 1989) and stable over a 9- to 10-year period (Heller et al., 2013; Christensen & Mendoza, 1986, Radler, Rigotti & Ryff, 2018). The subset of 31 subjects that completed this measure again 7 years later show an intraclass correlation of 0.797 between their two PWB scores. Thus, PWB scores likely reflect stable individual differences in well-being that can be linked with daily diary and fMRI data collected months or years later.

Daily Diary Procedure

The daily diary protocol included telephone interviews each evening for eight consecutive evenings. Participants reported on stressful events and a range of positive and negative emotions since the time they woke up. The frequency with which participants experienced each emotion that day was rated on a five-point scale (0 = none of the time, 4 = all of the time). Thirteen PA items were rated: “in good spirits,” “cheerful,” “extremely happy,” “calm and peaceful,” “satisfied,” “full of life,” “close to others,” “like you belong,” “enthusiastic,” “attentive,” “proud,” “active,” and “confident”; and fourteen NA items were rated: “restless or fidgety,” “nervous,” “worthless,” “so sad nothing cheered you up,” “everything required effort,” “hopeless,” “lonely,” “afraid,” “jittery,” “irritable,” “ashamed,” “upset,” “frustrated,” and “angry”. Cronbach’s alpha for the PA scale = 0.96 and for the NA scale = 0.91, based on reliability calculations recommended by Raudenbush and colleagues (Charles et al., 2016, 2019; Raudenbush et al., 1991). These items were averaged to create a daily PA and NA score. The average across all of a participant’s daily interviews was taken to reflect their mean PA and NA.

fMRI Procedure and Acquisition

Participants completed an image-viewing task in which they saw 60 positive, 60 negative, and 60 neutral images from the International Affective Picture System (Figure 1; Lang, Bradley, & Cuthbert, 2008). These images were matched for luminosity, picture complexity (using the jpeg file size as an index of complexity) and social content (van Reekum et al., 2018), and the positive and negative picture sets were equally arousing on average. Following a 1 second fixation screen, each image was presented for 4 seconds, and participants indicated with a manual response whether the image was positive, negative or neutral in valence. Following IAPS slides was either a neutral facial expression (1 or 3 seconds following the IAPS image offset), or a black screen with a fixation cross but no face. Faces were one of 30 male faces from the XM2VTSDB multi-modal face database (Messer et al., 1999) presented for 0.5 seconds (Figure 1). Participants were asked to simply view the faces and made no response to these stimuli. The total intertrial interval (including the face presentations for trials where faces followed an IAPS picture) was selected from an exponential distribution and varied from 5.5 to 17.6 seconds with an average duration = 8.89 seconds, and consisted of a black screen with a white fixation cross.

A standard clinical whole-head quadrature head coil was used. Five functional runs were collected. Functional images were obtained using a T2*-weighted, echo-planar images (EPI; [30 sagittal slices, 4 mm thickness with 1 mm gap; 3.75 x 3.75 mm in-plane (64 x 64 voxels); FOV = 240; repetition time (TR)/echo time (TE)/Flip, 2000 ms/30 ms/60°; 262 whole-brain volumes per run]. Additionally, a high-resolution T1-weighted anatomical image was obtained (T1-weighted inversion recovery fast gradient echo; 256 x 256 in-plane resolution; 240 mm FOV; 124 x 1.1 mm axial slices).

fMRI Analysis

fMRI data were resampled to 2 x 2 x 2mm then preprocessed and analyzed using AFNI (Cox, 1996), ANTs (Avants et al., 2011), and FSL (Jenkinson et al., 2012). T1 images were skull-stripped using FSL's BET function. For functional data, the first four volumes of each run were discarded before analysis and then functional images were despiked using the 3dDespike program from the AFNI toolbox. Motion correction was performed with the ANTs toolbox; functional images were first aligned within each run to the mean image, then all runs were aligned to the mean image from the first functional run. Next, using ANTs, nonlinear spatial normalization was applied to the functional data to match the MNI152 template. The normalized functional images were then brain-masked and scaled using FSL.

The preprocessed data were input to AFNI's 3dDeconvolve to test a general linear model (GLM) with task and nuisance regressors. The 5 functional runs were concatenated, and the model was estimated using these concatenated data. There were 9 task regressors: 3 regressors for the IAPs (one for each of the positive, negative, neutral conditions; convolved for 4 seconds), 3 regressors for neutral facial expressions 1 second after an image (one for each of the positive, negative, neutral images it follows; convolved for 0.5 seconds), and three for neutral facial expressions 3 seconds later (again one for each of the positive, negative, neutral images it follows; convolved for 0.5 seconds). The 6 standard motion regressors were included, as well as their derivatives, and the square of each of those 12 parameters, resulting in 24 motion parameters in the model to remove sources of variance due to motion (Satterthwaite et al., 2013). At the individual level, this GLM yielded whole-brain maps of beta coefficients for each voxel for each task condition.

Representational Similarity Analysis.

We used RSA to measure individual differences in the persistence of amygdala activity patterns in response to different classes of stimuli: negative images and neutral facial expression that follow negative images. In essence, this analysis quantifies the how closely the representation of neutral face expressions that are presented after negative stimuli resemble the representation of actual negative images in the amygdala. To accomplish this, the subject-level, whole-brain beta maps from the GLM were inputted to the RSA MATLAB toolbox (Nili et al., 2014). Beta maps were masked to isolate voxels of the left and right amygdala thresholded at 50% using the Harvard Oxford Subcortical Brain Atlas (Desikan et al., 2006). Given evidence from fMRI studies for amygdala laterality in some processing of affective stimuli (Baas et al., 2004; Dyck et al., 2011; Murphy et al., 2020), we examined the left and right amygdala separately. The 2 x 2 x 2mm left amygdala mask contained 240 voxels and the right amygdala mask contained 280 voxels. These masks correspond to approximately 27 and 32 voxels respectively in the native sampling space. The persistence of left and right amygdala activity from affective images to subsequent neutral facial expressions was quantified in the following way: we extracted the pattern of left and right amygdala activity in response to the IAPS image as well as in response to the neutral faces. Both patterns were reshaped into a $1 \times n$ vector and correlated with one another. These correlation values were Fisher Z transformed then subtracted from 1 to calculate the representational distance, or dissimilarity, between the two patterns. This typical dissimilarity metric yields values which range from 0 to 2, with greater numbers reflecting greater dissimilarity. Because this analysis was focused on persistence (a measure of similarity rather than dissimilarity), we inverted the dissimilarity metrics from the RSA analysis. Also, we elected to use the neutral facial expressions presented at 3s post-IAPS offset, rather than 1s, for our metric of amygdala persistence given the sluggish nature of the BOLD signal (Lindquist et al., 2009).

For the primary analyses, this persistence metric was calculated by including all stimuli from all runs into a single GLM. However, because within-run temporal autocorrelation could bias similarity metrics, we performed additional analyses in which each fMRI run was analyzed separately. This alternative method reduces potential biases introduced by temporal collinearity among events within the same run. Once each run was modeled with a separate GLM, this similarity metric was calculated by individually estimating the dissimilarity among all between-run, image-face comparisons (e.g., similarity of amygdala IAPS activity of run 1 with amygdala face activity in run 2, run 3, etc.). Once dissimilarity metrics for all 20 IAPS-face permutations were calculated they were averaged together. Importantly, this average value excludes the within-run comparisons (e.g., there is no comparison among run 1 IAPS slides with run 1 faces). Replicating our analyses using this potentially less biased metric demonstrates that the primary effects are unlikely to be a product of biased estimates driven by within-run temporal correlation.

Associations Between Amygdala Persistence and Daily Affect

After assessing zero-order correlations, we tested whether amygdala persistence following negative images was associated with mean daily PA and NA in multiple linear regressions. These regressions included covariates of age, gender, race, twin status, time between visits, and number of telephone diary interviews completed. Specifically, dummy-coded variables were included for race categories and each twin pair. We included persistence following neutral and positive images as covariates in all models to confirm that any effects were specific to amygdala persistence to negative images. We conducted an identical analysis, but with patterns of activity in the left and right occipital pole, the right and left nucleus accumbens (NAcc), and the medial prefrontal cortex (mPFC) to test whether effects were specific to the amygdala. Similar to the amygdala,

these regions were taken from the Harvard Oxford Brain Atlas and thresholded at 50% (Desikan et al., 2006). We selected the occipital pole because we did not expect persistence in these visual regions to be associated with daily affect. However, we selected the NAcc and the mPFC to explore whether other regions that are often implicated in affective processing show similar effects as the amygdala. False discovery rate (FDR; Benjamini & Hochberg, 1995) correction was applied to these correlation matrices by using the raw p-value vector as input to the *p.adjust* command in the *stats* package (i.e. `p.adjust(uncorrected_ps, method = "fdr")`). *p.adjust* also requires an argument for the number of unique pairwise comparisons of interest. In the 9 by 9 correlation matrix shown in Table 1, for example, the number of tests is 36 or the length of the raw p-value vector. Further, p-values obtained from multiple regression models for relationships involving variables of interest (daily NA, PA, PWB, and amygdala persistence) were concatenated into a single vector and FDR-corrected using *p.adjust*.

To test whether the RSA-derived persistence was superior to a univariate measure of persistence in predicting affect, we computed the difference between univariate amygdala responses to the IAPs images and the neutral facial expressions. Although it has been argued that within-person difference scores in neural activity may be unreliable as individual difference measures (Infantolino et al., 2018), this metric was most analogous to the RSA-derived persistence metric. We calculated mean beta value across the voxels in the left and right amygdala for negative, neutral, and positive IAPs images and also for those faces that followed those images. We then took the absolute value of the difference between the IAPs images and the face stimuli for each valence condition to obtain scores analogous to similarity, or representational distance. These univariate differences scores were then examined as predictors of daily affect and PWB in correlations and multiple regression models similar to the multivariate persistence metrics.

Finally, path analyses were conducted using version 0.6-4 of the 'lavaan' package in R (Rosseel, 2012) to test whether amygdala persistence predicted PWB, via daily affect. First, with the full sample, we assessed whether daily affect mediated the relationship between amygdala persistence and the first measurement of PWB. While this model takes advantage of all available data, the outcome precedes the predictors in time. To overcome this challenge, we also tested a similar model with the subset of participants who had completed a measurement of PWB after their fMRI and diaries and used this second measurement of PWB as the outcome. Further, while traditional mediation frameworks required that a significant total effect be present for amygdala persistence and PWB (Baron & Kenny, 1986), current methodological models of mediation suggest that this direct pathway is not necessary in order to detect mediation (Hayes, 2009). The significance of the path coefficients in these models was determined by using the Wald z-statistic, which divides the parameter estimate by its standard error value. The path model also included the identical covariates as above.

Results

Descriptive Statistics

The mean of the composite PWB score was 39.87 ($SD = 5.70$, range = 22.17 – 48.50). For the 31 participants that reported on PWB approximately 7 years later, the mean was 40.33 ($SD = 4.63$, range = 28.75 – 48.50). Participants completed an average of 7.5 daily telephone interviews ($SD = 0.87$, range = 4 – 8). The mean daily NA was 0.15 ($SD = 0.17$, range = 0 - 0.76); mean daily PA was 2.89 ($SD = 0.62$, range = 1.35 - 3.99). Daily affect variables were assessed for acceptable skew and kurtosis based on recommendations by Kline (i.e. skewness < |2| and kurtosis < |7|; Kline, 2015) with the

'psych' R package, version 1.8.12 (Revelle, 2019). While mean PA demonstrated acceptable skew (-0.23) and kurtosis (-0.22), mean NA showed a positive skew that approached unacceptable (1.89) with acceptable kurtosis (3.34). As a result, mean NA was square root transformed (transformed skewness = 0.59 and kurtosis = 0.25). Results from all analyses remained unchanged when tested with the raw NA values. Descriptive statistics for all variables are shown in Table 1.

Left amygdala persistence following a negative image ranged from 0.03 to 1.97, with a mean of 0.88 ($SD = 0.43$). Left amygdala persistence activity following a neutral image ranged from 0.06 to 1.78, with a mean of 0.80 ($SD = 0.41$). Left amygdala persistence following a positive image ranged from 0.01 to 1.74, with a mean of 0.84 ($SD = 0.42$).

Right amygdala persistence following a negative image ranged from 0.00 to 1.79 ($M = 0.71$, $SD = 0.44$). Right amygdala persistence following a neutral image ranged from 0.10 to 2.14 ($M = 0.79$, $SD = 0.40$) and right amygdala persistence following a positive image ranged from 0.01 to 1.97 ($M = 0.89$, $SD = 0.44$). A repeated measures ANOVA assessed differences in persistence by hemisphere (left and right) across the three image valence conditions. There was no main effect of hemisphere ($F(1,51) = 0.80$, $p = 0.372$) or valence $F(2,102) = 0.97$, $p = 0.382$, nor was there an interaction between hemisphere and valence $F(2,102) = 1.80$, $p = 0.167$).

Similarity in left and right amygdala persistence was also observed in Pearson bivariate correlations. There were positive correlations between left and right amygdala persistence following a negative image ($r = 0.36$, $p = 0.030$), a neutral image ($r = 0.36$, $p = 0.030$), and a positive image ($r = 0.45$, $p = 0.003$).

Links between Amygdala Persistence and Daily Affect

Pearson bivariate correlations (Table 1) revealed that left amygdala persistence following negative images was associated with daily mean NA ($r = 0.37, p = 0.040$; Figure 2A) and daily mean PA ($r = -0.47, p < 0.001$; Figure 2B). Left amygdala persistence following neutral images was not associated with daily mean NA ($r = 0.12, p = 0.626$) or PA ($r = -0.11, p = 0.645$). Similarly, persistence following positive images was unrelated to mean NA ($r = -0.16, p = 0.540$) or PA ($r = -0.06, p = 0.810$). In contrast with the left amygdala, right amygdala persistence following negative images was not significantly related to daily mean NA ($r = 0.21, p = 0.370$) or daily mean PA ($r = -0.12, p = 0.626$). The relationship between amygdala persistence and daily NA was not significantly different between the left and right amygdala ($t(51) = 1.04, p = 0.305$) in a test of difference between two correlated correlations in the *psych* package in R (version 1.8.12, Revelle, 2018; Stieger, 1980). However, the relationship between amygdala persistence and daily PA was significantly stronger in the left amygdala compared with the right ($t(51) = -2.45, p = 0.018$).

Given the associations between left amygdala persistence following negative images and daily affect, we tested the robustness of these effects in a multiple regression model that included relevant covariates. The effect of left amygdala persistence following the negative images on daily affect held when including covariates in multiple regressions. Greater amygdala persistence predicted more daily NA ($b = 0.24, t(40) = 3.46, p = 0.003, B = 0.50$) and less daily PA ($b = -0.97, t(40) = -4.84, p < 0.001, B = -0.67$) when controlling for age, gender, race, twin status, the time between visits, and the number of diaries completed. The relationship between left amygdala persistence following negative images and daily NA was also specific, as persistence following neutral images ($b = 0.07, t(38) = 1.00, p = 0.448, B = 0.014$), or following positive images ($b = -0.03, t(38) = -0.41, p = 0.730, B = -0.06$), did not attenuate the relationship between left amygdala persistence following negative images and daily NA

($b = 0.23$, $t(38) = 3.19$, $p = 0.006$, $B = 0.48$) when included in the model (adjusted $R^2 = 0.17$; Table 2). Similarly, left amygdala persistence following neutral images ($b = -0.08$, $t(38) = -0.40$, $p = 0.730$, $B = -0.06$), or following positive images ($b = -0.11$, $t(38) = -0.56$, $p = 0.700$, $B = -0.08$), did not attenuate the effect of persistence following negative images on daily PA ($b = -0.97$, $t(38) = -4.69$, $p < 0.001$, $B = -0.67$; full model adjusted $R^2 = 0.25$; Table 2).

Links between Amygdala Persistence and Psychological Well-being

Pearson bivariate correlations (Table 1) revealed that left amygdala persistence following negative images was not directly associated with PWB ($r = -0.16$, $p = 0.540$). Similarly, right amygdala persistence following negative images was not associated with PWB ($r = 0.05$, $p = 0.810$). Left amygdala persistence following neutral images was also not associated with PWB ($r = -0.03$, $p = 0.874$), nor was left amygdala persistence following positive images related to PWB ($r = -0.14$, $p = 0.600$). Further, right amygdala persistence following neutral images was not significantly associated with PWB ($r = -0.18$, $p = 0.480$), nor was right amygdala persistence following positive images related to PWB ($r = -0.03$, $p = 0.868$).

Links between Daily Affect and Psychological Well-being

Pearson bivariate correlations (Table 1) showed that PWB was significantly associated with daily mean PA ($r = 0.51$, $p < 0.001$), but not mean NA ($r = -0.18$, $p = 0.480$).

Daily Positive Affect links Left Amygdala Persistence following negative images to PWB.

Because there was not a significant direct association between PWB and daily mean NA, we only included daily mean PA in the full path models. In the first model

using the full sample, and in line with the bivariate correlations reported above, there was no direct effect of left amygdala persistence following negative images on the first PWB measurement ($b = 1.75$, $z = 0.84$, $p = 0.517$; Table 3). However, there were significant effects of both left amygdala persistence on mean PA ($b = -0.97$, $z = -5.49$, $p < 0.001$), and mean PA on PWB ($b = 4.83$, $z = 3.79$, $p < 0.001$). Furthermore, there was a significant indirect effect of left amygdala persistence on PWB through mean PA ($b = -4.68$, $z = -3.12$, $p = 0.005$; Figure 3). This model accounted for 43% of the variance in mean PA and 44% of the variance in the first measurement of PWB.

We also tested whether the indirect effect was also present using the second PWB measurement approximately seven years later as the outcome in a subset of 31 participants. This model overcomes the temporal limitations of using the first measurement of PWB, which was collected prior to the fMRI and daily diaries. A similar pattern of effects was found when repeating the model. There were significant effects of both left amygdala persistence on mean PA ($b = -0.88$, $z = -3.29$, $p = 0.003$), and mean PA on PWB ($b = 5.64$, $z = 4.34$, $p < 0.001$) yielding a significant indirect effect of left amygdala persistence on PWB through mean PA ($b = -4.94$, $z = -2.62$, $p = 0.016$). Unlike the previous model, there was a direct effect of left amygdala persistence following negative images on the second PWB measurement ($b = 4.72$, $z = 2.19$, $p = 0.047$). This model accounted for 46% of the variance in mean PA and 57% of the variance in the first measurement of PWB, suggesting that amygdala persistence is indirectly related to future PWB years later through daily PA.

Univariate Amygdala Difference Scores between Images and Neutral Facial Expressions

Overall, we did not find that univariate differences scores between IAPS images and subsequent faces were related to daily affect or PWB. Specifically, the left amygdala difference scores between negative images and the neutral facial expressions that

followed were not significantly related to daily NA ($r = 0.03$, $p = 0.94$), daily PA ($r = 0.03$, $p = 0.94$), or PWB ($r = -0.27$, $p = 0.20$). Left amygdala difference scores between the neutral IAPs images and the faces that follow were also not associated with daily NA ($r = 0.09$, $p = 0.79$), daily PA ($r = 0.08$, $p = 0.84$), or PWB ($r = -0.18$, $p = 0.46$). Further, left amygdala difference scores between the positive IAPs images and the subsequent faces were not significantly related to daily NA ($r = -0.01$, $p = 0.96$), daily PA ($r = 0.20$, $p = 0.42$), or PWB ($r = 0.12$, $p = 0.66$). Similarly, univariate difference scores in the right amygdala between negative images and the neutral facial expressions were not associated with daily NA ($r = 0.05$, $p = 0.94$), daily PA ($r = 0.02$, $p = 0.96$), or PWB ($r = -0.16$, $p = 0.52$). Additionally, right amygdala difference scores between neutral images and the faces that followed were not related to daily NA ($r = 0.01$, $p = 0.96$), daily PA ($r = 0.04$, $p = 0.94$), or PWB ($r = -0.22$, $p = 0.36$). Finally, right amygdala difference scores between the positive IAPs images and the subsequent faces were not significantly related to daily NA ($r = -0.05$, $p = 0.94$), daily PA ($r = 0.13$, $p = 0.66$), or PWB ($r = 0.05$, $p = 0.94$).

Similar to the RSA persistence metric, we tested whether the left amygdala univariate difference scores would predict individual differences in daily affect using multiple regression with relevant covariates. The univariate difference between negative images and neutral facial expressions that followed did not significantly predict daily mean NA ($b = 0.04$, $t(38) = 0.21$, $p = 0.839$, $B = 0.04$) in a model that also included the univariate difference between neutral images and neutral facial expressions that followed ($b = -0.02$, $t(38) = -0.11$, $p = 0.912$, $B = -0.02$), the univariate difference between positive images and neutral facial expressions that followed ($b = -0.007$, $t(38) = -0.03$, $p = 0.973$, $B = -0.005$), and the same covariates included in the other analyses (age, gender, race, twin status, the time between visits, and the number of diaries completed). A similar lack of effect resulted from an identical model with daily mean PA

as the outcome. The univariate difference between negative images and neutral facial expressions that followed did not significantly predict daily mean PA ($b = 0.40$, $t(38) = 0.63$, $p = 0.534$, $B = 0.12$), nor did the difference between neutral images and neutral facial expressions that followed ($b = -0.11$, $t(38) = -0.18$, $p = 0.861$, $B = -0.03$), or the difference between positive images and neutral facial expressions that followed ($b = 0.60$, $t(38) = 0.91$, $p = 0.370$, $B = 0.16$). Together, these results suggest that our RSA persistence metric captured unique individual variance related to daily affective experience.

Exploratory Analyses Addressing Spatial Specificity

To assess whether the association between daily affect and amygdala persistence following negative images was unique to the amygdala, we tested whether these effects were present in other regions of the brain. To do this, we extracted the patterns of the left and right occipital pole (LOcc; ROcc), the left and right NAcc (LNAcc; RNAcc), and the mPFC, then conducted an identical RSA analysis using these regions. The Pearson bivariate correlations are displayed in Table 4.

Results revealed that persistence in the LOcc and ROcc following negative images were not associated with daily mean NA (ROcc $r = 0.07$, $p = 0.837$; LOcc $r = 0.09$, $p = 0.810$), or mean PA (ROcc $r = -0.11$, $p = 0.774$; LOcc $r = -0.18$, $p = 0.667$), or PWB (ROcc $r = 0.12$, $p = 0.774$; LOcc $r = -0.05$, $p = 0.872$). Occipital persistence following neutral images was also not associated with daily mean NA (ROcc $r = -0.14$, $p = 0.765$; LOcc $r = -0.23$, $p = 0.495$), mean PA (ROcc $r = -0.05$, $p = 0.872$; LOcc $r = 0.07$, $p = 0.837$), or PWB (ROcc $r = 0.04$, $p = 0.873$; LOcc $r = 0.20$, $p = 0.650$). Also, occipital persistence following positive images was unrelated to daily mean NA (ROcc $r = -0.17$, $p = 0.671$; LOcc $r = -0.07$, $p = 0.837$), mean PA (ROcc $r = 0.24$, $p = 0.495$; LOcc $r = 0.11$, $p = 0.774$), or PWB (ROcc $r = 0.20$, $p = 0.644$; LOcc $r = 0.29$, $p = 0.322$).

Similarly, mPFC persistence following negative images was unrelated to daily NA ($r = 0.09$, $p = 0.810$), daily PA ($r = -0.11$, $p = 0.774$), or PWB ($r = -0.07$, $p = 0.837$). Also, mPFC persistence following neutral images was not significantly associated with daily NA ($r = 0.04$, $p = 0.873$), daily PA ($r = 0.03$, $p = 0.896$), or PWB ($r = 0.04$, $p = 0.873$) nor was mPFC persistence following positive images related to daily NA ($r = -0.01$, $p = 0.970$), daily PA ($r = -0.14$, $p = 0.765$), or PWB ($r = -0.25$, $p = 0.400$).

Results also showed that RNACC persistence following negative images was not associated with daily NA ($r = -0.16$, $p = 0.700$), daily PA ($r = 0.10$, $p = 0.810$), or PWB ($r = -0.10$, $p = 0.782$). Persistence following neutral images in the RNACC was similarly not related to daily NA ($r = -0.17$, $p = 0.668$), daily PA ($r = -0.05$, $p = 0.873$), or PWB ($r = -0.20$, $p = 0.644$). Further, RNACC persistence following positive images was unrelated to daily NA ($r = 0.12$, $p = 0.774$), daily PA ($r = -0.02$, $p = 0.912$), or PWB ($r = -0.23$, $p = 0.495$). We found that LNAcc persistence following negative images was significantly correlated with daily PA ($r = -0.51$, $p < 0.001$) as well as PWB ($r = -0.57$, $p < 0.001$), but not with daily NA ($r = 0.27$, $p = 0.364$). However, LNAcc persistence following neutral images was not related to any outcomes, including daily NA ($r = 0.21$, $p = 0.612$), daily PA ($r = -0.25$, $p = 0.400$), and PWB ($r = -0.09$, $p = 0.810$). LNAcc persistence following positive image was also not significantly associated with daily NA ($r = 0.04$, $p = 0.873$), daily PA ($r = 0.09$, $p = 0.810$), and PWB ($r = 0.02$, $p = 0.941$). These correlations suggest that greater persistence (i.e. similarity between representations of negative images and faces that follow after) in the LNAcc is also related to less PA in daily life and even less PWB.

These outcomes indicate that individual differences in persistence following negative images in the amygdala, but not cortical regions, is linked to daily affect. Additionally, the LNAcc, a subcortical structure that can mediate both appetitive and

defensive behaviors (Al-Hasani et al., 2015; Berridge, 2019), shows some overlap with the left amygdala in that is also correlated with daily PA, but not NA.

Exploratory Analyses to Reduce Impact of Temporal Autocorrelation of fMRI

Because within-run temporal autocorrelation of fMRI could impact metrics of similarity, we recomputed left amygdala persistence by calculating IAPS – face similarity permutations separately for each scan run and then averaged them together. This metric reflects the mean of all possible between-run image-face comparisons computed separately for negative, neutral, and positive image comparisons with the neutral facial expressions presented 3s later.

Links between left amygdala persistence and daily affect were similar, though slightly weaker for daily NA. Specifically, the relationship between this alternative metric of persistence and daily NA ($b = 0.56$, $t(36) = 2.00$, $p = 0.053$) was not quite significant. However, this alternative metric of persistence following negative images was associated with daily PA ($b = -2.82$, $t(36) = -3.22$, $p = 0.003$) in identical models from the primary analysis, which included age, gender, race, twin status, the time between visits, and the number of diaries completed as covariates. Also, similar to the primary analyses, this alternative metric of persistence following negative image was not directly associated with PWB ($b = 10.45$, $t(37) = 1.52$, $p = 0.128$), but the indirect effect of left amygdala persistence on PWB through daily PA was replicated using this alternative metric of persistence ($b = -12.95$, $z = -2.76$, $p = 0.006$).

Discussion

What gives rise to individual differences in both transient emotional experiences and enduring PWB across months and years? The field of translational neuroscience

asserts that real-world, affective outcomes can be predicted by neural activity observed in controlled, experimental settings (Berkman & Falk, 2013; Shackman et al., 2016). In particular, individual differences in neural activity in regions encoding value and salience, like the amygdala, may be critical. Previous research has shown that more persistent amygdala activity in response to aversive stimuli predicts higher levels of traits including neuroticism (Schuyler et al., 2014). We extended this literature by examining how experimentally-obtained individual differences in amygdala persistence, quantified using RSA, map onto naturalistic daily affect and PWB assessed over long periods of time.

Individuals with greater left amygdala persistence following negative images reported more NA and less PA in daily life. This aligns with previous research demonstrating that the amygdala facilitates rapid, negative appraisals of neutral stimuli preceded by an unrelated, aversive stimulus (Grupe et al., 2018; Lapate et al., 2016; Tambini et al., 2017). Additionally, previous work documenting individual differences in amygdala persistence following emotional stimuli (Schuyler et al., 2014; Siegle et al., 2002) suggests that some individuals may be more susceptible to the biasing of neutral stimuli by unrelated, negative stimuli. In the current study, we cannot confirm the neutral facial expressions were perceived as neutral, however, individual differences in neutral-face perception were accounted for by including neutral- and positive-image persistence as covariates. These results demonstrate that amygdala similarity between negative images and neutral facial expressions that follow such images is linked to subjective emotional experience outside of the laboratory. It may be that for individuals with greater amygdala persistence, negative moments may become amplified or prolonged by imbuing unrelated moments that follow with a negative appraisal. Ultimately, this persistence could result in more NA and less PA, on average, in daily life. This brain-behavior link between left amygdala persistence and daily affect can inform our understanding of more enduring, long-term evaluations of well-being.

PWB was indirectly linked to left amygdala persistence through daily affect. Those who exhibited less amygdala persistence, reported higher daily PA, which in turn, was associated with more positive PWB evaluations. A significant indirect effect between two variables in the absence of a direct effect can occur for a myriad of reasons (Hayes, 2009; Pardo & Román, 2013; MacKinnon et al., 2000; Shrout & Bolger, 2002). One reason is the length of time between variable measurements (Pardo & Román, 2013; James et al., 2006). It is possible that the extended time interval separating amygdala persistence and PWB measurements (mean = 41 months) limited our ability to detect a significant direct effect. Another possibility is that amygdala persistence contributes to momentary stimulus appraisals, which are proximal to day-to-day affect. In contrast, PWB is determined by an adjoining set of complex integrative processes that govern how one recollects, integrates, and judges the meaning of these daily experiences. These results build upon previous links between daily PA and greater PWB (Burns & Ma, 2015) by tying variability in complex, cognitive evaluations about well-being to the emotional landscape of daily life and individual differences in brain function. Critically, this indirect effect held regardless of whether we examined the first measurement of PWB in the full sample or the second measurement of PWB 7 years after the fMRI collected in a subset of participants. In our study, lower daily NA was not significantly associated with PWB. It is difficult to know how negative moments are integrated into broader evaluations of one's well-being, but one possibility is that less variability in mean NA relative to mean PA in this sample limited our ability to explain individual differences in NA. Another possibility is that the null result between daily NA and PWB may be due a lack of power, as previous research with over 2,000 subjects (Rush et al., 2019) found a link between daily NA and PWB. Therefore, additional research should probe the directionality and role of daily NA in linking individual differences in brain function with PWB.

Our persistence results were found in the left amygdala, indicating potential differential affective processing between the left and right amygdala. However, hemispheric differences should be interpreted with caution as only the correlation between amygdala persistence and PA (not NA) was significantly stronger in left amygdala. Still, the left amygdala has been reportedly more involved with cognitive processing of emotional information compared to the right amygdala, which facilitates automatic stimulus detection and processing (Dyck et al., 2011; Gläscher & Adolphs, 2003). Thus, left amygdala persistence may reflect more elaborative, cognitive appraisal of the negative stimuli (Phelps et al., 2001). Another possibility is that the left-lateralized amygdala findings may be an artifact of data preprocessing procedures, such as smoothing and motion correction (Murphy et al., 2020). Given the inconsistencies in the current study and the broader literature, additional research is needed to clarify the potential lateralization in amygdala persistence.

These findings also extend the previous literature by using RSA to quantify amygdala persistence in relationship to psychological functioning. We applied RSA, a multivariate technique (Kriegeskorte et al., 2008; Brooks & Freeman, 2018), to capture individual differences in neural persistence, or the discriminability of neural representations of temporally related stimulus classes. In other words, we quantified, for a given person, how much the voxel-wise representation of negative image stimuli resembled the representation of neutral facial expressions that followed negative stimuli in the amygdala. Previous research on amygdala persistence that averages across amygdala voxels ignores differential activation patterns across the spatially heterogeneous amygdala structure. Indeed, we found that univariate left amygdala responses to neutral facial expressions following negative images were not associated with daily NA or PA. However, our persistence metric, derived from an RSA comparing voxel-wise patterns of activity in the amygdala between affective stimuli and neutral

facial expressions that followed, did predict daily NA and PA. These dynamic comparisons better match real-world experiences than isolating response to a single stimulus, making them prime targets for quantifying amygdala persistence.

These results suggest numerous avenues for future research. Specifically, additional task features could provide a more comprehensive account of neural persistence. The post-IAPs image stimuli used here were limited to neutral facial expressions, however we cannot be certain that subjects perceived these stimuli as neutral. Because the amygdala is involved in facial processing (Todorov et al., 2008), the social nature of these stimuli may have influenced our metric of persistence. It is important to examine whether other types of putatively neutral stimuli influence amygdala persistence and to collect concurrent subjective ratings of these stimuli. Additionally, including greater variation in the time between affective and neutral stimuli (i.e. 3, 5 and 7 seconds) could pinpoint the optimal lag for detecting individual variation in persistence. Although we used individuals' neutral and positive persistence values as covariates, greater exploration of individual factors that influence amygdala persistence, such as attentional state and hemodynamic response function shape, could deepen our understanding of the mechanisms of neural persistence. Additional work should examine persistence both in the presence and the absence of subsequent, intervening stimuli to better understand influences on enduring emotional representations over time.

Beyond the fMRI task design, additional research should further probe the mechanisms by which amygdala persistence translates to daily affect by exploring the cortical-subcortical relationships that explain individual differences in amygdala persistence. Moreover, the amygdala persistence effects here should be replicated with higher spatial resolution scans. In this study, the voxel size was relatively large and sampling amygdala activity with higher spatial resolution would increase the number of voxels included yielding a richer analysis. Beyond the amygdala, other brain regions

should be further investigated. Although persistence in cortical regions did not show significant relationships with daily affect and PWB, left NAcc persistence was negatively related to daily PA, similar to the left amygdala. This suggests that persistence-affect relationships may not be specific to the amygdala. Also, while this subset of the MIDUS sample captured a range of adults (ages 39-76), it would be useful to examine these effects during periods of critical development, such as adolescence (Steinberg, 2005). Finally, although daily diary protocols often map on well to momentary experience (Kahneman et al., 2004), ecological momentary assessments across a longer timescale might provide a more fine-grained measurement of daily affective experiences.

In summary, individual differences in amygdala persistence were linked to meaningful, ecological affective outcomes. Specifically, the results indicated that those with less persistence in the left amygdala following negative images may be less susceptible to the biasing of neutral facial expression by unrelated, negative stimuli. Protection against this vulnerability to negative biasing may bring about more PA as well as less NA in daily life. Critically, it is through greater daily PA that left amygdala persistence is related to PWB. As PWB is a complex, cognitive self-evaluation requiring one to integrate across a lifetime of experiences, such an evaluative process is likely supported by distributed brain networks rather than any one individual region.

References

- Al-Hasani, R., McCall, J. G., Shin, G., Gomez, A. M., Schmitz, G. P., Bernardi, J. M., Pyo, C.-O., Park, S. I., Marcinkiewicz, C. M., Crowley, N. A., Krashes, M. J., Lowell, B. B., Kash, T. L., Rogers, J. A., & Bruchas, M. R. (2015). Distinct Subpopulations of Nucleus Accumbens Dynorphin Neurons Drive Aversion and Reward. *Neuron*, 87(5), 1063–1077. <https://doi.org/10.1016/j.neuron.2015.08.019>
- Avants, B. B., Tustison, N. J., Song, G., Cook, P. A., Klein, A., & Gee, J. C. (2011). A reproducible evaluation of ANTs similarity metric performance in brain image registration. *NeuroImage*, 54(3), 2033–2044. <https://doi.org/10.1016/j.neuroimage.2010.09.025>
- Baas, D., Aleman, A., & Kahn, R. S. (2004). Lateralization of amygdala activation: A systematic review of functional neuroimaging studies. *Brain Research Reviews*, 45(2), 96–103. <https://doi.org/10.1016/j.brainresrev.2004.02.004>
- Baron, R. M., & Kenny, D. A. (1986). The moderator–mediator variable distinction in social psychological research: Conceptual, strategic, and statistical considerations. *Journal of Personality and Social Psychology*, 51(6), 1173–1182. <https://doi.org/10.1037/0022-3514.51.6.1173>
- Benjamini, Y., & Hochberg, Y. (1995). Controlling the False Discovery Rate: A Practical and Powerful Approach to Multiple Testing. *Journal of the Royal Statistical Society. Series B (Methodological)*, 57(1), 289–300. JSTOR.
- Berkman, E. T., & Falk, E. B. (2013). Beyond Brain Mapping: Using Neural Measures to Predict Real-World Outcomes. *Current Directions in Psychological Science*, 22(1), 45–50. <https://doi.org/10.1177/0963721412469394>

- Berridge, K. C. (2019). Affective valence in the brain: Modules or modes? *Nature Reviews Neuroscience*, 20(4), 225–234. <https://doi.org/10.1038/s41583-019-0122-8>
- Brooks, J. A., & Freeman, J. B. (2018). Conceptual knowledge predicts the representational structure of facial emotion perception. *Nature Human Behaviour*, 2(8), 581–591. <https://doi.org/10.1038/s41562-018-0376-6>
- Burns, R. A., & Ma, J. (2015). Examining the association between psychological wellbeing with daily and intra-individual variation in subjective wellbeing. *Personality and Individual Differences*, 82, 34–39. <https://doi.org/10.1016/j.paid.2015.02.023>
- Charles, S. T., Mogle, J., Leger, K. A., & Almeida, D. M. (2019). Age and the Factor Structure of Emotional Experience in Adulthood. *The Journals of Gerontology: Series B*, 74(3), 419–429. <https://doi.org/10.1093/geronb/gbx116>
- Charles, S. T., Mogle, J., Urban, E. J., & Almeida, D. M. (2016). Daily Events are Important for Age Differences in Mean and Duration for Negative Affect but not Positive Affect. *Psychology and Aging*, 31(7), 661–671. <https://doi.org/10.1037/pag0000118>
- Cox, R. W. (1996). AFNI: Software for Analysis and Visualization of Functional Magnetic Resonance Neuroimages. *Computers and Biomedical Research*, 29(3), 162–173. <https://doi.org/10.1006/cbmr.1996.0014>
- Christensen, L. & Mendoza, J.L. (1986). A method of assessing change in a single subject: An alteration of the RC index. *Behavior Therapy*, 17, 305–308

- Davis, M., & Whalen, P. J. (2001). The amygdala: Vigilance and emotion. *Molecular Psychiatry*, 6(1), 13–34. <https://doi.org/10.1038/sj.mp.4000812>
- Desikan, R. S., Ségonne, F., Fischl, B., Quinn, B. T., Dickerson, B. C., Blacker, D., Buckner, R. L., Dale, A. M., Maguire, R. P., Hyman, B. T., Albert, M. S., & Killiany, R. J. (2006). An automated labeling system for subdividing the human cerebral cortex on MRI scans into gyral based regions of interest. *NeuroImage*, 31(3), 968–980. <https://doi.org/10.1016/j.neuroimage.2006.01.021>
- Dyck, M., Loughhead, J., Kellermann, T., Boers, F., Gur, R. C., & Mathiak, K. (2011). Cognitive versus automatic mechanisms of mood induction differentially activate left and right amygdala. *NeuroImage*, 54(3), 2503–2513. <https://doi.org/10.1016/j.neuroimage.2010.10.013>
- Gläscher, J., & Adolphs, R. (2003). Processing of the Arousal of Subliminal and Supraliminal Emotional Stimuli by the Human Amygdala. *Journal of Neuroscience*, 23(32), 10274–10282. <https://doi.org/10.1523/JNEUROSCI.23-32-10274.2003>
- Grupe, D. W., Schaefer, S. M., Lapate, R. C., Schoen, A. J., Gresham, L. K., Mumford, J. A., & Davidson, R. J. (2018). Behavioral and neural indices of affective coloring for neutral social stimuli. *Social Cognitive and Affective Neuroscience*, 13(3), 310–320. <https://doi.org/10.1093/scan/nsy011>
- Hayes, A. F. (2009). Beyond Baron and Kenny: Statistical mediation analysis in the new millennium. *Communication Monographs*, 76(4), 408–420.

- Heller, A. S., van Reekum, C. M., Schaefer, S. M., Lapate, R. C., Radler, B. T., Ryff, C. D., & Davidson, R. J. (2013). Sustained striatal activity predicts eudaimonic well-being and cortisol output. *Psychological science*, 24(11), 2191-2200
- Infantolino, Z. P., Luking, K. R., Sauder, C. L., Curtin, J. J., & Hajcak, G. (2018). Robust is not necessarily reliable: From within-subjects fMRI contrasts to between-subjects comparisons. *NeuroImage*, 173, 146–152.
<https://doi.org/10.1016/j.neuroimage.2018.02.024>
- James, L. R., Mulaik, S. A., & Brett, J. M. (2006). A tale of two methods. *Organizational Research Methods*, 9, 233-244.
- Jenkinson, M., Beckmann, C. F., Behrens, T. E. J., Woolrich, M. W., & Smith, S. M. (2012). FSL. *NeuroImage*, 62(2), 782–790.
<https://doi.org/10.1016/j.neuroimage.2011.09.015>
- Kahneman, D., Krueger, A. B., Schkade, D. A., Schwarz, N., & Stone, A. A. (2004). A Survey Method for Characterizing Daily Life Experience: The Day Reconstruction Method. *Science*, 306(5702), 1776–1780.
<https://doi.org/10.1126/science.1103572>
- Kline, R. B. (2015). *Principles and practice of structural equation modeling*. Guilford publications.
- Kober, H., Barrett, L. F., Joseph, J., Bliss-Moreau, E., Lindquist, K., & Wager, T. D. (2008). Functional grouping and cortical–subcortical interactions in emotion: A meta-analysis of neuroimaging studies. *NeuroImage*, 42(2), 998–1031.
<https://doi.org/10.1016/j.neuroimage.2008.03.059>

- Kriegeskorte, N., & Bandettini, P. (2007). Analyzing for information, not activation, to exploit high-resolution fMRI. *NeuroImage*, 38(4), 649–662.
<https://doi.org/10.1016/j.neuroimage.2007.02.022>
- Kriegeskorte, N., Mur, M., Ruff, D. A., Kiani, R., Bodurka, J., Esteky, H., Tanaka, K., & Bandettini, P. A. (2008). Matching Categorical Object Representations in Inferior Temporal Cortex of Man and Monkey. *Neuron*, 60(6), 1126–1141.
<https://doi.org/10.1016/j.neuron.2008.10.043>
- Lang, P. J., Bradley, M. M., & Cuthbert, B. N. (2008). International Affective Picture System (IAPS): Affective Ratings of Pictures and Instruction Manual. Technical Report A-8. Gainesville, FL: University of Florida.
- Lapate, R. C., Rokers, B., Tromp, D. P. M., Orfali, N. S., Oler, J. A., Doran, S. T., Adluru, N., Alexander, A. L., & Davidson, R. J. (2016). Awareness of Emotional Stimuli Determines the Behavioral Consequences of Amygdala Activation and Amygdala-Prefrontal Connectivity. *Scientific Reports*, 6.
<https://doi.org/10.1038/srep25826>
- Lapate, Regina C., Samaha, J., Rokers, B., Hamzah, H., Postle, B. R., & Davidson, R. J. (2017). Inhibition of Lateral Prefrontal Cortex Produces Emotionally Biased First Impressions: A Transcranial Magnetic Stimulation and Electroencephalography Study. *Psychological Science*, 28(7), 942–953.
<https://doi.org/10.1177/0956797617699837>
- LeDoux, J. E. (1996). *The emotional brain: The mysterious underpinnings of emotional life*. Simon & Schuster.
- LeDoux, J. E. (2000). Emotion Circuits in the Brain. *Annual Review of Neuroscience*, 23(1), 155–184. <https://doi.org/10.1146/annurev.neuro.23.1.155>

- Lindquist, M. A., Meng Loh, J., Atlas, L. Y., & Wager, T. D. (2009). Modeling the hemodynamic response function in fMRI: Efficiency, bias and mis-modeling. *NeuroImage*, 45(1, Supplement 1), S187–S198.
<https://doi.org/10.1016/j.neuroimage.2008.10.065>
- MacKinnon, D. P., Krull, J. L., & Lockwood, C. M. (2000). Equivalence of the mediation, confounding and suppression effect. *Prevention Science*, 1(4), 173–181.
- Messer, K., Matas, J., Kittler, J., Luettin, J., & Maitre, G. (1999). *XM2VTSDB: The Extended M2VTS Database*.
<http://www.ee.surrey.ac.uk/Research/VSSP/xm2vtsdb>
- Murphy, J. E., Yanes, J. A., Kirby, L. A. J., Reid, M. A., & Robinson, J. L. (2020). Left, right, or bilateral amygdala activation? How effects of smoothing and motion correction on ultra-high field, high-resolution functional magnetic resonance imaging (fMRI) data alter inferences. *Neuroscience Research*, 150, 51–59.
<https://doi.org/10.1016/j.neures.2019.01.009>
- Nili, H., Wingfield, C., Walther, A., Su, L., Marslen-Wilson, W., & Kriegeskorte, N. (2014). A Toolbox for Representational Similarity Analysis. *PLOS Computational Biology*, 10(4), e1003553. <https://doi.org/10.1371/journal.pcbi.1003553>
- Ölander, K., Muukkonen, I., Numminen, J., & Salmela, V. (2017). Representational similarity analysis of EEG and fMRI responses to face identities and emotional expressions. *Journal of Vision*, 17(10), 271–271.
<https://doi.org/10.1167/17.10.271>
- Pardo, A., & Román, M. (2013). Reflections on the Baron and Kenny model of statistical mediation. *Anales de Psicología*, 29(2), 614–623.

- Phelps, E. A., O'Connor, K. J., Gatenby, J. C., Gore, J. C., Grillon, C., & Davis, M. (2001). Activation of the left amygdala to a cognitive representation of fear. *Nature Neuroscience*, 4(4), 437–441. <https://doi.org/10.1038/86110>
- Radler, B. T., Rigotti, A., & Ryff, C. D. (2018). Persistently high psychological well-being predicts better HDL cholesterol and triglyceride levels: Findings from the midlife in the U.S. (MIDUS) longitudinal study. *Lipids in Health and Disease*, 17(1), 1. <https://doi.org/10.1186/s12944-017-0646-8>
- Raudenbush, S. W., Rowan, B., & Kang, S. J. (1991). A Multilevel, Multivariate Model for Studying School Climate with Estimation Via the EM Algorithm and Application to U. S. High-School Data. *Journal of Educational Statistics*, 16(4), 295–330. JSTOR. <https://doi.org/10.2307/1165105>
- Revelle, W. (2018) psych: Procedures for Personality and Psychological Research, Northwestern University, Evanston, Illinois, USA. R Package version 1.8.12, <https://CRAN.R-project.org/package=psych>.
- Rosseel, Y. (2012). lavaan: An R Package for Structural Equation Modeling. *Journal of Statistical Software*, 48(1), 1–36. <https://doi.org/10.18637/jss.v048.i02>
- Rush, J., Rast, P., Almeida, D. M., & Hofer, S. M. (2019). Modeling long-term changes in daily within-person associations: An application of multilevel SEM. *Psychology and aging*, 34(2), 163–176. <https://doi.org/10.1037/pag0000331>
- Ryff, C. D. (1989). Happiness is everything, or is it? Explorations on the meaning of psychological well-being. *Journal of Personality and Social Psychology*, 57(6), 1069–1081. <https://doi.org/10.1037/0022-3514.57.6.1069>

- Ryff, C. D., & Keyes, C. L. M. (1995). The structure of psychological well-being revisited. *Journal of Personality and Social Psychology*, 69(4), 719–727. <https://doi.org/10.1037/0022-3514.69.4.719>
- Sander, D., Grafman, J., & Zalla, T. (2003). The Human Amygdala: An Evolved System for Relevance Detection. *Reviews in the Neurosciences*, 14(4). <https://doi.org/10.1515/REVNEURO.2003.14.4.303>
- Satterthwaite, T. D., Elliott, M. A., Gerraty, R. T., Ruparel, K., Loughead, J., Calkins, M. E., Eickhoff, S. B., Hakonarson, H., Gur, R. C., Gur, R. E., & Wolf, D. H. (2013). An Improved Framework for Confound Regression and Filtering for Control of Motion Artifact in the Preprocessing of Resting-State Functional Connectivity Data. *NeuroImage*, 64. <https://doi.org/10.1016/j.neuroimage.2012.08.052>
- Schuyler, B. S., Kral, T. R. A., Jacquart, J., Burghy, C. A., Weng, H. Y., Perlman, D. M., Bachhuber, D. R. W., Rosenkranz, M. A., MacCoon, D. G., van Reekum, C. M., Lutz, A., & Davidson, R. J. (2014). Temporal dynamics of emotional responding: Amygdala recovery predicts emotional traits. *Social Cognitive and Affective Neuroscience*, 9(2), 176–181. <https://doi.org/10.1093/scan/nss131>
- Shackman, A. J., Tromp, D. P. M., Stockbridge, M. D., Kaplan, C. M., Tillman, R. M., & Fox, A. S. (2016). Dispositional negativity: An integrative psychological and neurobiological perspective. *Psychological Bulletin*, 142(12), 1275–1314. <https://doi.org/10.1037/bul0000073>
- Shrout, P. E., & Bolger, N. (2002). Mediation in experimental and nonexperimental studies: New procedures and recommendations. *Psychological Methods*, 7(4), 422–445. <https://doi.org/10.1037/1082-989X.7.4.422>

- Siegle, G. J., Steinhauer, S. R., Thase, M. E., Stenger, V. A., & Carter, C. S. (2002). Can't shake that feeling: Event-related fMRI assessment of sustained amygdala activity in response to emotional information in depressed individuals. *Biological Psychiatry*, 51(9), 693–707. [https://doi.org/10.1016/S0006-3223\(02\)01314-8](https://doi.org/10.1016/S0006-3223(02)01314-8)
- Steinberg, L. (2005). Cognitive and affective development in adolescence. *Trends in Cognitive Sciences*, 9(2), 69–74. <https://doi.org/10.1016/j.tics.2004.12.005>
- Steiger, J. H. (1980). Tests for comparing elements of a correlation matrix. *Psychological Bulletin*, 87, 245-251.
- Tambini, A., Rimmele, U., Phelps, E. A., & Davachi, L. (2017). Emotional brain states carry over and enhance future memory formation. *Nature Neuroscience*, 20(2), 271–278. <https://doi.org/10.1038/nn.4468>
- Todorov, A., Baron, S. G., & Oosterhof, N. N. (2008). Evaluating face trustworthiness: A model based approach. *Social Cognitive and Affective Neuroscience*, 3(2), 119–127. <https://doi.org/10.1093/scan/nsn009>
- van Reekum, C. M., Schaefer, S. M., Lapate, R. C., Norris, C. J., Tun, P. A., Lachman, M. E., Ryff, C. A., & Davidson, R. J. (2018). Aging is associated with a prefrontal lateral-medial shift during picture-induced negative affect. *Social Cognitive and Affective Neuroscience*, 13(2), 156–163. <https://doi.org/10.1093/scan/nsx144>

Figure and Table Legends

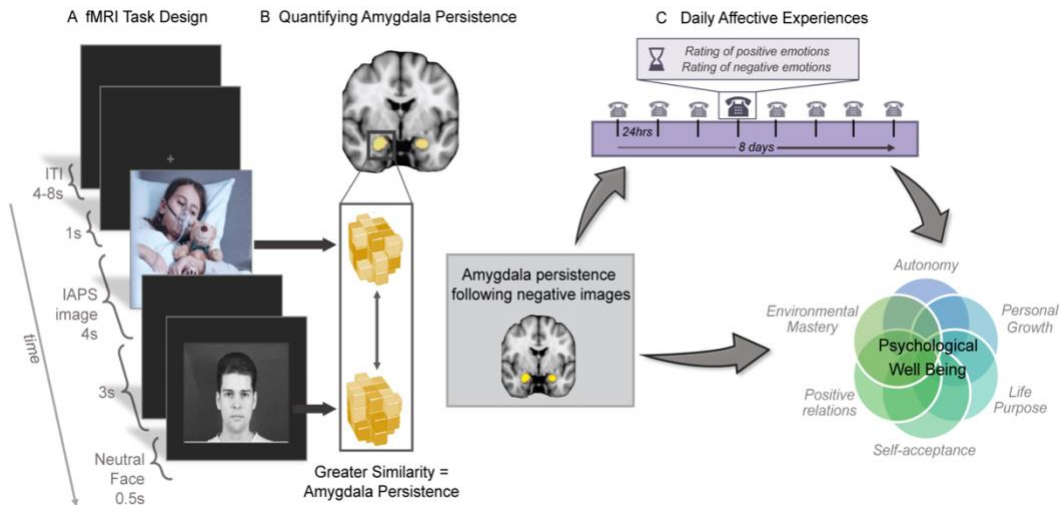


Figure 1. Study Design Schematic. (A) fMRI Task Design: Participants viewed 60 negative, neutral, and positive IAPS images for 4 seconds. When viewing the image, participants indicated whether the image was negative, neutral, or positive via a button press. Images were followed by either a neutral facial expression displayed for 0.5 seconds or a black screen. On neutral facial expression trials, the face was displayed 3 seconds after the offset of the IAPS image. We note that there were 20 trials with faces displayed 1 second after the image, which are not pictured here, that were not analyzed in relation to the affective outcomes of interest. Images here are examples of images seen not actual images used. (B) Quantifying Amygdala Persistence: a representational similarity analysis was conducted to calculate the voxel-wise pattern similarity in the amygdala between the negative images and the neutral faces that followed them. (C) Daily Affective Experience: participants reported on the degree to which they experienced 13 positive and 14 negative emotions throughout their day. This information was collected over the telephone to an interviewer for 8 consecutive days. Individual differences in daily affective experience was used as a mediator linking the metric of amygdala persistence with psychological well-being.

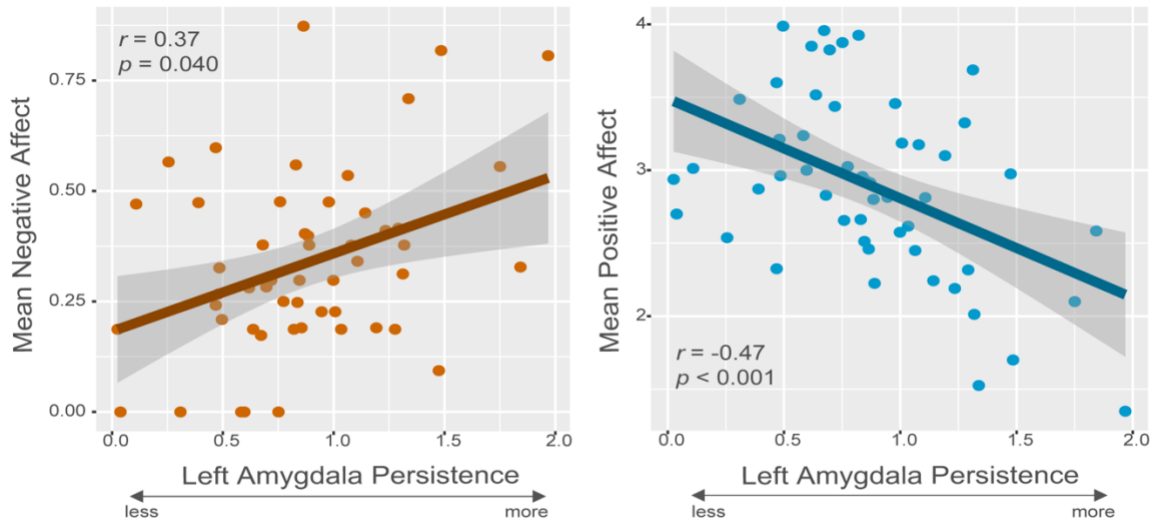


Figure 2. Left amygdala persistence following negative images is associated with daily negative and positive affect. Zero-order Pearson correlation plots between left amygdala persistence following negative images and mean NA and PA from the daily diaries. The square root transformed mean negative affect is shown here. $N = 52$.

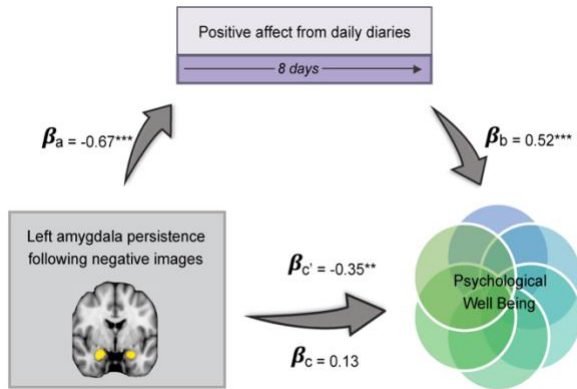


Figure 3. Path model of left amygdala persistence following negative images predicting psychological well-being via daily positive affect. Standardized path coefficients displayed. Covariates, including age, gender, number of diaries completed, race, twin status, time between assessments and persistence following neutral and positive images are not shown here for simplicity. The indirect path is denoted with c'. Significant paths denoted with bold font. $N = 52$. * $p < .05$; ** $p < .01$; *** $p < .001$.

Table 1. Means, standard deviations, and Pearson bivariate correlation coefficients between affective variables and amygdala persistence.

Variable	<i>M (SD)</i>	1	2	3	4	5	6	7	8
1. PWB	39.87 (5.70)								
2. Mean Daily PA	2.89 (0.62)	.51***							
3. Mean Daily NA	0.34 (0.20)	-0.18	-0.66***						
4. LAmyg. Persistence: Negative images	0.88 (0.38)	-0.16	-.47**	.37*					
5. LAmyg..Persistence: Neutral images	0.81 (0.36)	-0.03	-0.11	0.12	0.04				
6. LAmyg. Persistence: Positive images	0.86 (0.37)	-0.14	-0.06	-0.16	-0.16	0.13			
7. RAmyg. Persistence: Negative images	0.74 (0.38)	0.05	-0.12	0.21	0.36*	0.30	0.23		
8. RAmyg. Persistence: Neutral images	0.80 (0.35)	-0.18	0.06	-0.35*	0.07	0.36*	0.40*	0.01	
9. RAmyg. Persistence: Positive images	0.91 (0.38)	-0.03	-0.13	0.06	0.10	-0.10	0.45**	0.05	0.23

Legend. *N* = 52. *M* and *SD* are used to represent mean and standard deviation, respectively. Mean Daily Negative Affect is square root transformed here. Correlations are FDR corrected. PWB = psychological well-being, PA = positive affect, NA = negative affect, Amyg. = Amygdala, L = left, R = Right. * *p* < .05; ** *p* < .01; *** *p* < .001.

Table 2. Multiple regression models of daily negative and positive affect predicted by left amygdala persistence following negative images

Predictor	<i>b</i> (<i>SE</i>)	β	<i>t</i>	<i>p</i> -value
Negative Affect Model:				
Left Amygdala Persistence: Negative Images	0.23 (0.07) **	0.48	3.20	0.006
Left Amygdala Persistence: Neutral Images	0.07 (0.07)	0.14	1.00	0.45
Left Amygdala Persistence: Positive Images	-0.03 (0.07)	-0.06	-0.41	0.73
Time between fMRI and Diaries	-0.00 (0.00)	0.29	1.22	0.23
Diary Days Completed	-0.07 (0.03)	-0.28	-1.89	0.07
Race	-0.04 (0.09)	-0.08	-0.39	0.70
Gender	-0.03 (0.06)	-0.08	-0.54	0.60
Age	-0.00 (0.00)	-0.07	-0.47	0.64
Positive Affect Model:				
Left Amygdala Persistence: Negative Images	-0.97 (0.21)***	-0.67	-4.69	< 0.001
Left Amygdala Persistence: Neutral Images	-0.08 (0.21)	-0.06	-0.40	0.73
Left Amygdala Persistence: Positive Images	-0.11 (0.20)	-0.08	-0.56	0.70
Time between fMRI and Diaries	-0.00 (0.00)	-0.11	-0.31	0.63
Diary Days Completed	0.02 (0.10)	0.02	0.15	0.88
Race	-0.23 (0.27)	-0.18	-0.86	0.40
Gender	0.13 (0.18)	0.10	0.69	0.50
Age	0.01 (0.00)	0.18	1.28	0.21

Legend. *df* = 37, * *p* < .05; ** *p* < .01; *** *p* < .001. Mean daily negative affect is square root transformed in this model. Five dummy-coded variables to represent the five sets of twins in the sample were included in the model and not shown here. Race was coded as 1 = 'Caucasian' and 0 = 'African American'. Gender was coded as 0 = 'Male' and 1 = 'Female'.

Table 3. Path model of left amygdala persistence following negative images on the first measure of psychological well-being through daily positive affect

Path	<i>b</i> (<i>SE</i>)	β	<i>Z</i>	95% Lower CI	95% Upper CI	<i>p</i> -value
a	-0.97 (0.18) ***	-0.67	-5.49	-1.32	-0.62	< 0.001
b	4.83 (1.28) ***	0.52	3.79	2.33	7.32	< 0.001
c	1.75 (2.09)	0.13	0.84	-2.34	5.84	0.517
c'	-4.68 (1.50) **	-0.35	-3.12	-7.63	-1.74	0.005
total	-2.93 (1.89)	-0.22	-1.55	-6.64	0.77	0.120

Legend. *N* = 52. Path *a* is the effect of left amygdala persistence on mean daily PA, path *b* is the effect of mean daily PA on PWB, and path *c* is the effect of left amygdala persistence on PWB. Path *c'* is the indirect effect of left amygdala persistence on PWB through mean daily PA. Covariates included in the model were age, gender, race, twin status, number of diaries completed, and time between assessments. * *p* < .05; ** *p* < .01; *** *p* < .001

Table 4. Pearson bivariate correlations between PWB, daily affect, and persistence within other brain regions

Variable	1	2	3
1. PWB			
2. Mean Daily PA	0.51**		
3. Mean Daily NA	-0.18	-0.66**	
4. L. Occ. Persistence: Negative images	-0.05	-0.18	0.09
5. L. Occ. Persistence: Neutral images	0.20	0.07	-0.23
6. L. Occ. Persistence: Positive images	0.29	0.11	-0.07
7. R. Occ. Persistence: Negative images	0.12	-0.11	0.07
8. R. Occ. Persistence: Neutral images	0.04	-0.05	-0.14
9. R. Occ. Persistence: Positive images	0.20	0.24	-0.17
10. L. NAcc. Persistence: Negative images	-0.57**	-0.51**	0.27
11. L. NAcc. Persistence: Neutral images	-0.09	-0.25	0.21
12. L. NAcc. Persistence: Positive images	0.02	0.09	0.04
13. R. NAcc. Persistence: Negative images	-0.10	0.10	-0.16
14. R. NAcc. Persistence: Neutral images	-0.20	-0.05	-0.17
15. R. NAcc. Persistence: Positive images	-0.23	-0.02	0.12
16. mPFC Persistence: Negative images	-0.07	-0.11	0.09
17. mPFC Persistence: Neutral images	0.04	0.03	0.04
18. mPFC Persistence: Positive images	-0.25	-0.14	-0.01

Legend. $N = 52$. Mean daily negative affect is square root transformed here. PWB = psychological well-being, PA = positive affect, NA = negative affect, R = right, L = left, Occ = Occipital, NAcc = Nucleus accumbens, mPFC = medial prefrontal cortex. * $p < .05$; ** $p < .01$; *** $p < .001$